

Recap

Binary hypothesis testing: P_0, P_1 distributions on \mathcal{X}

$$T: \mathcal{X}^n \rightarrow \{0, 1\}$$

False positive: $\mathbb{P}_{\bar{x} \sim P_0^n} [T(\bar{x}) = 1] = \alpha(T)$

False negative: $\mathbb{P}_{\bar{x} \sim P_1^n} [T(\bar{x}) = 0] = \beta(T)$

$$\left. \begin{array}{l} \alpha(T) \\ \beta(T) \end{array} \right\} \alpha(T) + \beta(T) \geq 1 - \delta_{\mathbb{T}_n}(P_0^n, P_1^n)$$

Ratio tests: $T(\bar{x}) = 1$ iff $\frac{P_1^n(\bar{x})}{P_0^n(\bar{x})} \geq \Delta$

\Leftrightarrow

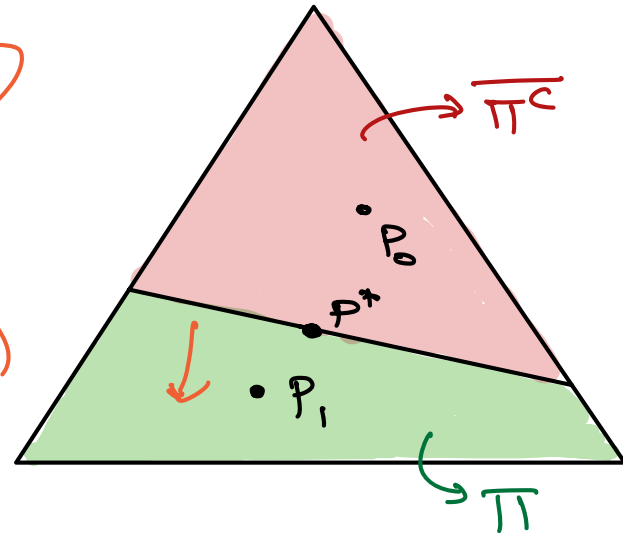
$$\underbrace{D(P_{\bar{x}} \parallel P_0) - D(P_{\bar{x}} \parallel P_1)} \geq \frac{1}{n} \log \Delta$$

$$P_{\bar{x}} \in \Pi$$

Behavior of exponents

$$\alpha(\tau) \approx 2^{-n \cdot D(P^* \parallel P_0)} \approx 2^{-n D(P_i \parallel P_0)}$$

$$\beta(\tau) \approx 2^{-n \cdot D(P^* \parallel P_i)} \approx 2^{-n D(P_0 \parallel P_i)}$$



depends on Δ

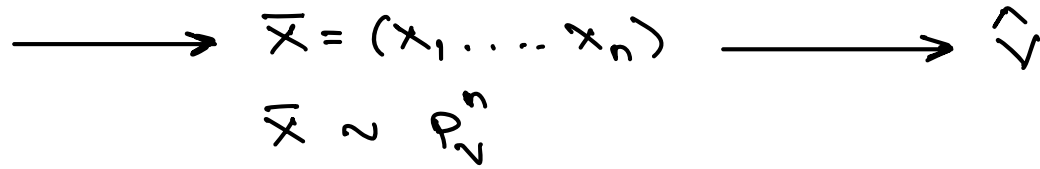
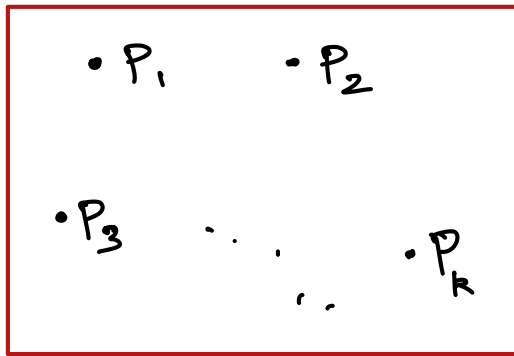
$$\frac{1}{10} \alpha(\tau) + \frac{9}{10} \beta(\tau) \approx 2^{-n \cdot D(P_0 \parallel P_i)} + 2^{-n \cdot D(P_i \parallel P_0)}$$

$$\approx 2^{-n \cdot \min\{D(P_0 \parallel P_i), D(P_i \parallel P_0)\}}$$

$C(P_0, P_i)$

$$\alpha(\tau) + \beta(\tau) \approx 1 - \frac{1}{2} \parallel \hat{P}_0 - \hat{P}_i \parallel_1 \approx 1 - \frac{1}{2} \sqrt{2 \ln 2 \cdot n D(P_0 \parallel P_i)}$$

Multiple hypotheses: Bound via Fano's inequality



$\mathcal{V} \equiv$ Collection of k hypotheses

$$P_e = \mathbb{P}[\mathcal{V} \neq \hat{\nu}]$$

a.v. ν : uniform on \mathcal{V}

$$\text{(Fano): } H_2(P_e) + P_e \log(|\mathcal{V}| - 1) \geq H(\nu | \bar{X})$$

$$\triangleright P_e \geq 1 - \frac{n \cdot \mathbb{E}_{\nu_1, \nu_2 \in \mathcal{V}} [D(P_{\nu_1} \| P_{\nu_2})] + 1}{\log |\mathcal{V}|}$$

Proof:

$$1 + p_e \cdot \log(2) \geq \underbrace{H_2(p_e)}_{\leq 1} + p_e \cdot \underbrace{\log(2)}_{\leq \log(2)} \geq H(V|X)$$

$\geq \log(2) = \frac{H(V)}{\log(2)} - I(V; X)$

$$I(V; \bar{X}) = D(P(V, \bar{X}) \parallel P(V) P(\bar{X}))$$

$$= \underbrace{D(P(V) \parallel P(V))}_{=0} + D(P(\bar{X}|V) \parallel P(\bar{X}))$$

↓

$$\begin{aligned} & \mathbb{E}_{V_1} D(P_{V_1} \parallel \mathbb{E}_{V_2} P_{V_2}) \\ & \leq \left(\mathbb{E}_{V_1, V_2} D(P_{V_1} \parallel P_{V_2}) \right) \cdot n \end{aligned}$$

Minimax rates

How many samples do we need to learn properties of distributions

$$\Theta : \underbrace{\Pi}_{\text{Set of distributions}} \rightarrow \underbrace{\mathcal{H}}_{\text{Set of properties}}$$

$$\hat{\Theta} : \mathcal{X}^n \rightarrow \mathcal{H}$$

estimator for Θ

$$l : \mathcal{H} \times \mathcal{H} \rightarrow \mathbb{R}_+$$

loss function

$$l(\Theta(P), \hat{\Theta}(X))$$

Minimax risk

$$M_n(\pi, \ell) = \inf_{\hat{\theta}} \sup_{P \in \pi} \mathbb{E}_{\bar{x} \sim P^n} \ell(\hat{\theta}(\bar{x}), \theta(P))$$

↖ ↗
best estimator worst distribution

e.g. $\Theta = \mathcal{V}$ $\pi = \{P_v\}_{v \in \mathcal{V}}$ $\ell(\theta, \hat{\theta}) = \begin{cases} 1 & \theta \neq \hat{\theta} \\ 0 & \theta = \hat{\theta} \end{cases}$

$\theta(P_v) = v$

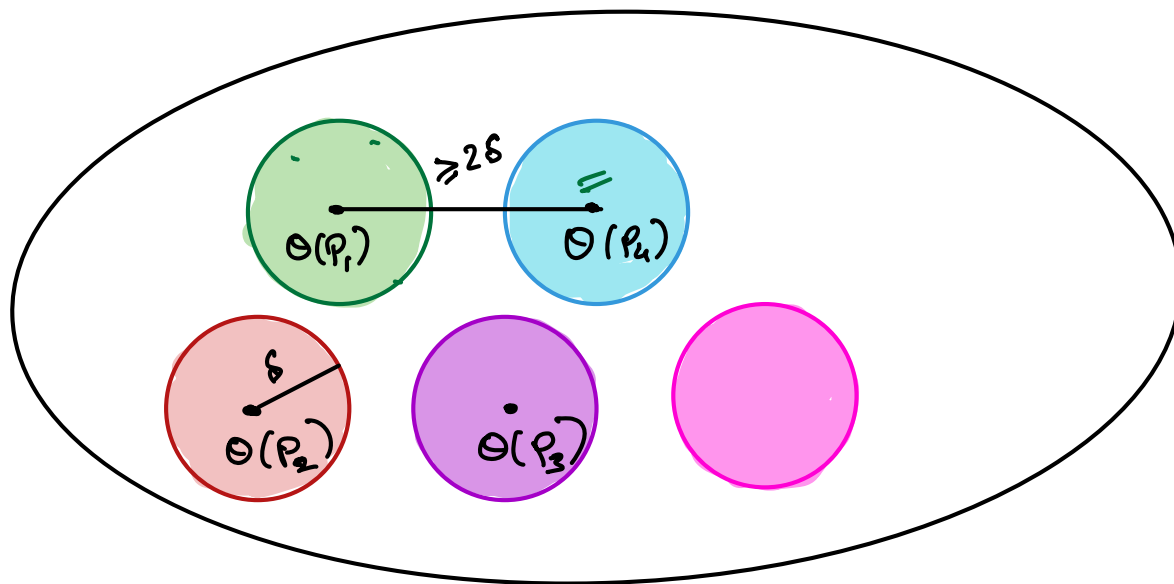
$$M_n(\pi, \ell) = \inf_{\hat{\theta}} \sup_{P_v} \mathbb{E}_{\bar{x} \sim P_v^n} \mathbb{P}[\hat{\theta}(\bar{x}) \neq v]$$

$$\geq \inf_{\hat{\theta}} \mathbb{E}_{v \in \mathcal{V}} \mathbb{E}_{\bar{x} \sim P_v^n} \mathbb{P}[\hat{\theta}(\bar{x}) \neq v]$$

~
 P_e
 for Fano

Minimax lower bounds via hypothesis testing

$$L(\theta, \hat{\theta}) = \underbrace{\Phi}_{\substack{\text{increasing} \\ \text{function}}}(\underbrace{\rho(\theta, \hat{\theta})}_{\text{distance metric}})$$



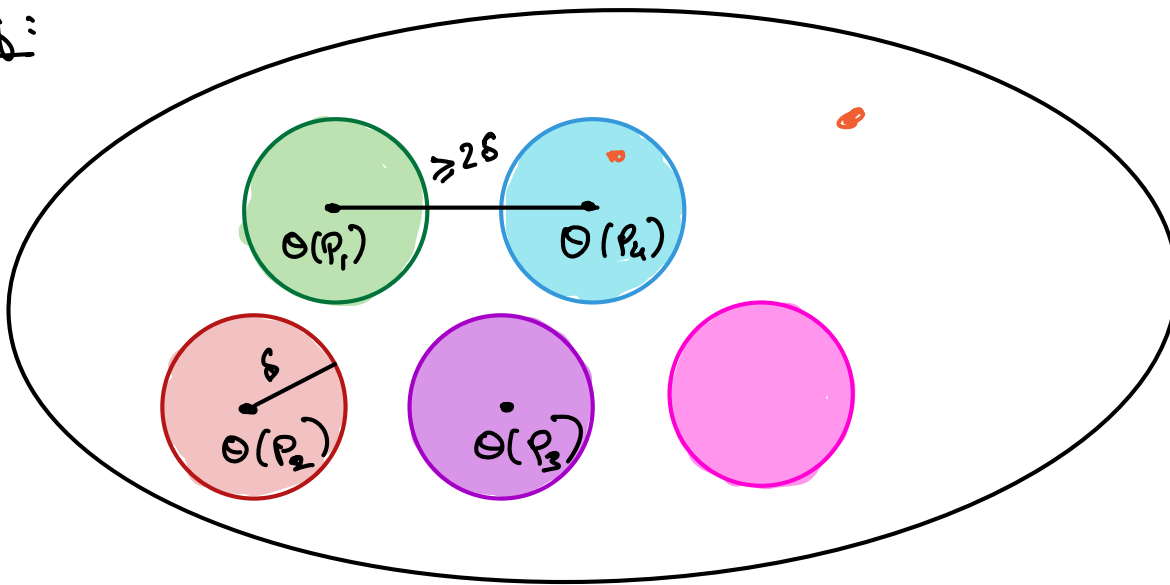
Cost of a mistake $\geq \Phi(\delta)$

Let $\{P_v\}_{v \in \mathcal{V}} \subseteq \Pi$ be a finite set of distributions s.t.

$$\rho(\theta(P_{v_1}), \theta(P_{v_2})) \geq 2\delta \quad \forall v_1 \neq v_2 \quad (\rho = \Phi(\rho(\theta, \hat{\theta})))$$

Then, $M_n(\Pi, \rho) \geq \Phi(\delta) \cdot \inf_T \left\{ \mathbb{E} \sum_{v \in \mathcal{V}} \mathbb{P}_{\mathbf{x} \sim P_v^n} [\tau(\mathbf{x}) \neq v] \right\}$
 $T: \mathcal{X}^n \rightarrow \mathcal{V}$

Proof:



Estimator $\hat{\theta}$
 can be converted
 to predictor T

$$T = \arg \min_{v \in \mathcal{V}} \rho(\theta(P_v), \hat{\theta})$$

$$\begin{aligned}
 M_n(\pi, \rho) &\stackrel{\text{if}}{\geq} \sup_{P \in \mathcal{P}_n} \mathbb{E}_{X \sim P} [\Phi(\rho(\theta(P)), \hat{\theta}(X))] \\
 &\geq \inf_{P \in \mathcal{P}_n} \mathbb{E}_{X \sim P} [\Phi(\rho(\theta(P)), \hat{\theta}(X))] \\
 &\geq \Phi(\delta) \cdot \mathbb{1}_{\{\tau(X) \neq v\}}
 \end{aligned}$$

One dimensional mean estimation

$$\mathcal{X} = \{0, 1\}$$

$\Pi =$ all distributions on \mathcal{X}

$$\theta(p) = \mathbb{E}_{x \sim p}[x] = p(1)$$

$$L(\theta, \hat{\theta}) = (\theta - \hat{\theta})^2 = \underbrace{(1\theta - \hat{\theta})^2}_p$$

Empirical estimator: $\hat{\theta}(\bar{x}) = \hat{\theta}(x_1, \dots, x_n) = \frac{1}{n} \sum_{i=1}^n x_i$

Say $\theta(p) = \mu$

$$E_{\bar{x} \sim p^n} [(\hat{\theta}(\bar{x}) - \mu)^2] = \frac{\mu(1-\mu)}{n}$$

Ex:

Lower bound (Le Cam's method)

$$P_0 = \begin{cases} 1 & \text{w.p. } 1/2 \\ 0 & \text{w.p. } 1/2 \end{cases}$$

$$P_1 = \begin{cases} 1 & \text{w.p. } 1/2 + 2\delta \\ 0 & \text{w.p. } 1/2 - 2\delta \end{cases}$$

$$(\mathcal{V} = \{0, 1\})$$

$$\sqrt{M}_n(\pi, \mathcal{L}) \geq \Phi(\delta) \cdot \inf_T \mathbb{E}_{\bar{x} \sim P_{\nu}^n} \mathbb{P}[\tau(\bar{x}) \neq \nu]$$

$$\geq \delta^2 \cdot \inf_T \left\{ \frac{1}{2} \mathbb{P}_{\bar{x} \sim P_0^n} [\tau(\bar{x}) = 1] + \frac{1}{2} \mathbb{P}_{\bar{x} \sim P_1^n} [\tau(\bar{x}) = 0] \right\}$$

$$\geq \delta^2 \cdot \frac{1}{2} \left(1 - \frac{1}{2} \sqrt{2 \ln 2 \cdot n \cdot D(P_0 \parallel P_1)} \right)$$

$\underbrace{\hspace{10em}}_{C \cdot \delta^2}$

Choosing δ

$$\sqrt{M_n(\pi, \mathcal{L})} \geq \frac{\delta^2}{2} \cdot \left(1 - \frac{1}{2} \|P_0^n - P_1^n\| \right)$$

$$\geq \frac{\delta^2}{2} \left(1 - \frac{1}{2} \cdot \sqrt{2 \ln 2 \cdot n \cdot C \cdot \delta^2} \right)$$

$$\frac{(C')^2}{2} \downarrow$$

$$\delta = C' \cdot \frac{1}{\sqrt{n}} \rightarrow \leq \frac{1}{2}$$

High-dimensional mean estimation

$$\Pi = \left\{ \mathcal{N}(\underbrace{\mu}_{\theta}, \mathbb{I}_d) \mid \mu \in \mathbb{R}^d \right\}$$

$$\ell(\theta, \hat{\theta}) = \|\theta - \hat{\theta}\|^2$$

Empirical estimator: $\hat{\theta}(\bar{x}) = \hat{\theta}(x_1 \dots x_n) = \frac{1}{n} \sum_{i=1}^n x_i$

$$\text{error} = \mathbb{E}_{\bar{x} \sim (\mathcal{N}(\mu, \mathbb{I}_d))^n} \left[\|\hat{\theta}(\bar{x}) - \mu\|^2 \right]$$

$$= \frac{d}{n}$$