

Gaussian Distributions (Recap)

"Standard Normal" $X \sim N(0, 1)$

density $\gamma(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$

$$\mathbb{E}[X] = 0, \quad \mathbb{E}[X^2] = \text{Var}[X] = 1, \quad \mathbb{E}[e^{\lambda X^2}] = \frac{1}{\sqrt{1-2\lambda}} \quad (\text{for } \lambda < \frac{1}{2})$$

Multiple independent $X_1, \dots, X_n \sim N(0, 1)$

- $C_1 X_1 + C_2 X_2 + \dots + C_n X_n + \mu \sim N(\mu, \|C\|^2)$

- $\gamma(x_1, \dots, x_n) = \frac{1}{(2\pi)^{n/2}} e^{-(x_1^2 + \dots + x_n^2)/2} = \frac{1}{(2\pi)^{n/2}} e^{-\|x\|^2/2}$

Johnson-Lindenstrauss dimension reduction [JL84]

► Let $P \subseteq \mathbb{R}^d$ be a set of points. $|P| = n$

$\forall \epsilon > 0 \quad (\epsilon < 1) \quad \exists$ linear transformation $\varphi: \mathbb{R}^d \rightarrow \mathbb{R}^k$

$$k \leq \frac{12 \ln n}{\epsilon^2 - \epsilon^3}$$

s.t. $\forall u, v \in P$

want
 $k \ll d$

$$(1-\epsilon) \cdot \|u-v\|^2 \leq \|\varphi(u) - \varphi(v)\|^2 \leq (1+\epsilon) \cdot \|u-v\|^2$$

Designing φ . (at random)

- Pick (random) matrix $G \in \mathbb{R}^{k \times d}$ s.t.

$$\forall i \in [k], j \in [d]. \quad G_{ij} \sim \mathcal{N}(0, 1)$$

& all entries are independent



- $\forall u \in \mathbb{R}^d$, define $\varphi(u) = \frac{Gu}{\sqrt{k}}$

Fix $u, v \in \mathcal{P}$. Suffices to prove

$$\mathbb{P}_G \left[(1-\epsilon) \|u-v\|^2 \leq \|\varphi(u) - \varphi(v)\|^2 \leq (1+\epsilon) \|u-v\|^2 \right] \geq 1 - \frac{2}{n^3}$$

$\hookrightarrow$ for $k = \frac{12 \ln n}{\epsilon^2 - \epsilon^3}$

$$\mathbb{P}_G \left[\forall u, v \in \mathcal{P} \quad (1-\epsilon) \|u-v\|^2 \leq \|\varphi(u) - \varphi(v)\|^2 \leq (1+\epsilon) \|u-v\|^2 \right]$$

$$\geq 1 - \binom{n}{2} \cdot \frac{2}{n^3}$$

$$\geq 1 - \frac{1}{n}$$

$$\mathbb{P} \left[(1-\epsilon) \|u-v\|^2 \leq \|\varphi(u) - \varphi(v)\|^2 \leq (1+\epsilon) \cdot \|u-v\|^2 \right]$$

• ||

$$\mathbb{P} \left[(1-\epsilon) \cdot \|u-v\|^2 \leq \frac{\|G(u-v)\|^2}{k} \leq (1+\epsilon) \cdot \|u-v\|^2 \right]$$

$$\mathbb{P} \left[(1-\epsilon) \cdot k \leq \frac{\|G(u-v)\|^2}{\|u-v\|^2} \leq (1+\epsilon) \cdot k \right] \quad w = \frac{u-v}{\|u-v\|}$$

$$\mathbb{P} \left[(1-\epsilon) \cdot k \leq \|G w\|^2 \leq (1+\epsilon) \cdot k \right]$$

Fix $\omega \in \mathbb{R}^d$, $\|\omega\| = 1$

Diagram illustrating the structure of a matrix G (size $k \times d$) and its multiplication with a vector w (size $d \times 1$).

The matrix G is shown as a grid with a red line indicating a row. The vector w is shown as a column.

The resulting vector is shown as a column with the first element highlighted in a blue box:

$$w_1 \cdot G_{11} + w_2 \cdot G_{12} + \dots + w_d \cdot G_{1d}$$

$$x_i = w_1 \cdot G_{i1} + \dots + w_d \cdot G_{id} \sim \mathcal{N}(0, \frac{\|w\|^2}{d})$$

$$\|Gv\|^2 = Z = X_1^2 + \dots + X_n^2$$

Sums of independent squared Gaussians

$$Z = X_1^2 + \dots + X_k^2$$

$$E[Z] = k$$

$X_1, \dots, X_k \sim N(0,1)$
independent

$$P[Z \geq (1+\epsilon) \cdot k] = P[e^{\lambda Z} \geq e^{\lambda(1+\epsilon)k}]$$

$$\boxed{\lambda < \frac{1}{2} \quad (\lambda > 0)}$$

$$\leq \frac{E[e^{\lambda Z}]}{e^{\lambda(1+\epsilon)k}} \quad e^{\lambda Z} = \prod_{i=1}^k e^{\lambda X_i^2}$$

$$= \frac{\prod_{i=1}^k E[e^{\lambda X_i^2}]}{e^{\lambda(1+\epsilon)k}} = \frac{?}{e^{\lambda(1+\epsilon)k}}$$

$$P[Z \geq (1+\epsilon) \cdot k] \leq \frac{\left(\frac{1}{\sqrt{1-2\lambda}}\right)^k}{e^{\lambda(1+\epsilon)k}} = \left(\frac{e^{-2\lambda(1+\epsilon)}}{1-2\lambda}\right)^{k/2}$$

$$\text{choose } \lambda = \frac{\epsilon}{2(1+\epsilon)}$$

$$\begin{aligned} \left(e^{-\epsilon}(1+\epsilon)\right)^{k/2} &\leq \left(\left(1-\epsilon + \frac{\epsilon^2}{2}\right)(1+\epsilon)\right)^{k/2} \\ &= \left(1-\epsilon^2 + \frac{\epsilon^2}{2} + \frac{\epsilon^3}{2}\right)^{k/2} \\ &\leq e^{-(\epsilon^2 - \epsilon^3)k/4} \end{aligned}$$

$$\underline{\text{Ex:}} \quad P[Z \leq (1-\epsilon)k] \leq e^{-(\epsilon^2 - \epsilon^3)k/4}$$

Putting things together

$$\mathbb{P} \left[\left| \underbrace{\|Gw\|^2}_Z - k \right| \geq \epsilon k \right] \leq 2 \cdot e^{-(\epsilon^2 - \epsilon^3) \cdot k/4}$$

$$\leq \frac{2}{n^3} \quad k \geq \frac{12 \ln n}{\epsilon^2 - \epsilon^3}$$

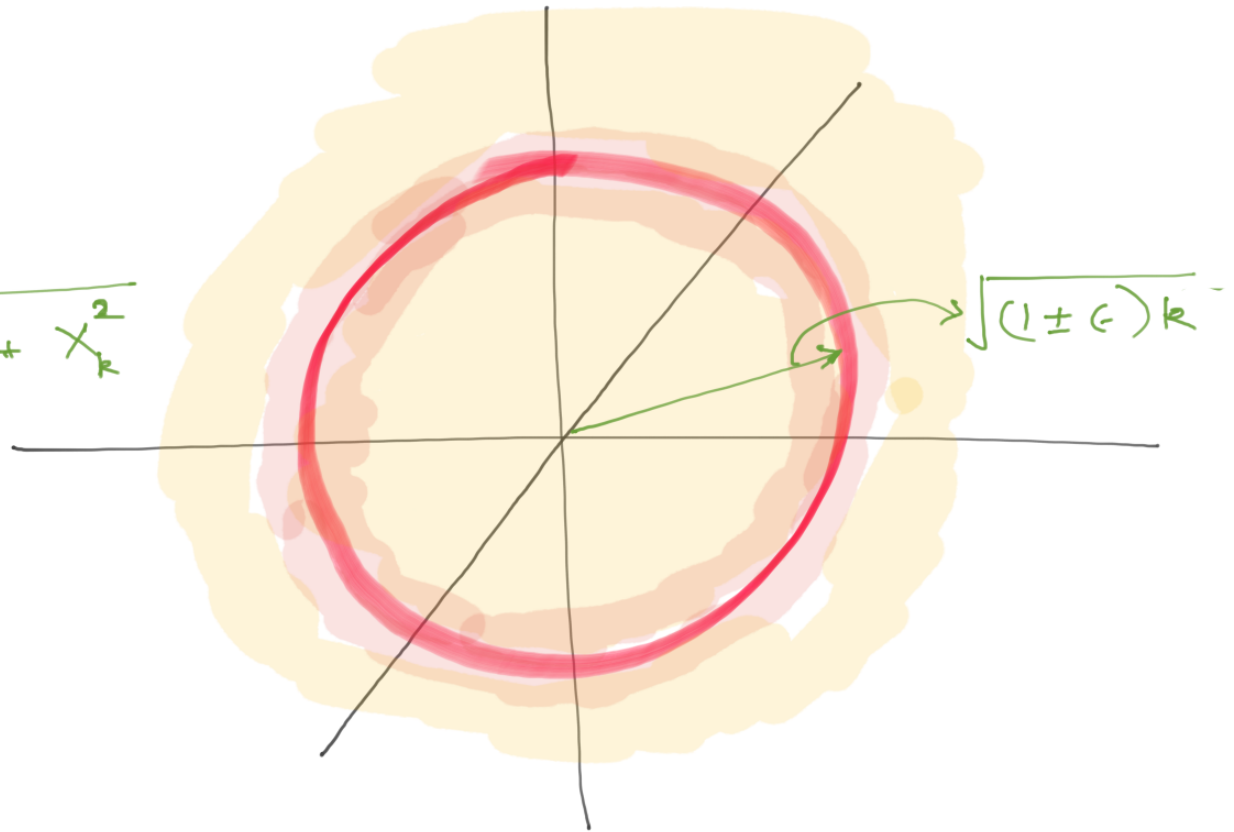
$$\mathbb{P} \left[\exists u, v \in \mathcal{P} . \left| \frac{\|G(u-v)\|^2}{k} - \|u-v\|^2 \right| \geq \epsilon \|u-v\|^2 \right] \leq$$

Multi-dimensional Spherical Gaussians

$$X = \begin{pmatrix} x_1 \\ \vdots \\ x_k \end{pmatrix}$$

$$\|X\| = \sqrt{x_1^2 + \dots + x_k^2}$$

$$\frac{e^{-\|x\|^2/2}}{(2\pi)^{k/2}}$$



Gordon's thm : infinite union bounds

$$T \subset \mathbb{R}^d \cap \{w \mid \|w\| = 1\}$$

$$\mathbb{P} \left[\exists w \in T, \left| \frac{\|Gw\|^2}{k} - 1 \right| \geq \epsilon \right] \leq \dots$$

$\gamma(T) \equiv$ Gaussian width

$$\mathbb{E} \sup_{x \sim \mathcal{N}(0,1)} | \langle x, w \rangle |$$

Dealing with general Gaussian vectors

Let $x_1, \dots, x_n \sim \mathcal{N}(0, 1)$ be independent.

$$\underline{y} = A\underline{x} + \underline{\mu}$$

$$\text{Let } \underline{y} = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = \begin{pmatrix} A \end{pmatrix} \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix} + \begin{pmatrix} \mu_1 \\ \vdots \\ \mu_n \end{pmatrix}$$

$$\underline{\mathbb{E}}[\underline{y}] = \underline{\mu}$$

mean (vector)

$$\underline{\mathbb{E}}[(\underline{y} - \underline{\mathbb{E}}[\underline{y}])(\underline{y} - \underline{\mathbb{E}}[\underline{y}])^T] = AA^T$$

covariance (matrix)

Dealing with general Gaussian vectors

Let $\underline{Y} = \begin{pmatrix} Y_1 \\ \vdots \\ Y_n \end{pmatrix}$ be a Gaussian random vector

with $E[\underline{Y}] = \underline{\mu}$ $E[(\underline{Y} - \underline{\mu})(\underline{Y} - \underline{\mu})^T] = \Sigma$

Then $\underline{Y}$ can be expressed as $\underline{Y} = \Sigma^{1/2} \underline{X} + \underline{\mu}$

for $\underline{X} = \begin{pmatrix} X_1 \\ \vdots \\ X_n \end{pmatrix} \sim \mathcal{N}(\underline{0}, \underline{I})$.

density $\gamma(\underline{Y}) = ?$

Diagonalizing Σ

Let $\Sigma = \sigma_1^2 \cdot \underbrace{w_1 w_1^*}_{\substack{\checkmark \\ \text{orthonormal basis} \\ \text{of eigenvectors}}} + \dots + \sigma_n^2 \cdot \underbrace{w_n w_n^*}_{\substack{\checkmark \\ \text{orthonormal basis} \\ \text{of eigenvectors}}}$

$$\begin{aligned} \Sigma^{-1/2} \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix} &= \begin{pmatrix} | & & | \\ w_1 & \dots & w_n \\ | & & | \end{pmatrix} \begin{pmatrix} \sigma_1 & & 0 \\ & \ddots & \\ 0 & & \sigma_n \end{pmatrix} \begin{pmatrix} -w_1^* \\ \vdots \\ -w_n^* \end{pmatrix} \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix} \\ &= \begin{pmatrix} | & & | \\ w_1 & \dots & w_n \\ | & & | \end{pmatrix} \begin{pmatrix} \sigma_1 & & 0 \\ & \ddots & \\ 0 & & \sigma_n \end{pmatrix} \begin{pmatrix} \langle w_1, x \rangle \rightarrow x'_1 \\ \vdots \\ \langle w_n, x \rangle \rightarrow x'_n \end{pmatrix} \end{aligned}$$

$$\begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix} = \begin{pmatrix} \hline \Sigma_1^* \hline \vdots \hline \vdots \hline \Sigma_n^* \hline \end{pmatrix} \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix} = \Sigma^* \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix}$$

$$\Sigma = \begin{pmatrix} \hline \Sigma_1 \hline \vdots \hline \vdots \hline \Sigma_n \hline \end{pmatrix} \begin{pmatrix} \sigma_1 & & 0 \\ & \ddots & \\ 0 & & \sigma_n \end{pmatrix} \begin{pmatrix} x_1' \\ \vdots \\ x_n' \end{pmatrix} = \Sigma \begin{pmatrix} \sigma_1 & & 0 \\ & \ddots & \\ 0 & & \sigma_n \end{pmatrix} x'$$

$$\mathbb{E}[X'] = 0$$

$$\mathbb{E}[(X' - \mathbb{E}X')(X' - \mathbb{E}X')] = I$$

Multi-dimensional Gaussians

